Mixture densities and the EM algorithm

- Mixture density with K components: $p(\mathbf{x}; \Theta) = \sum_{k=1}^{K} p(\mathbf{x}|k) p(k) \begin{cases} p(\mathbf{x}|k) & \text{component densities} \\ p(k) = \pi_k & \text{mixture proportions.} \end{cases}$
- Ex: Gaussian mixture: $\mathbf{x}|k \sim \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$. Mixture parameters: $\boldsymbol{\Theta} = \{\pi_k, \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k\}_{k=1}^K$.
- Maximum likelihood estimation of Gaussian mixture parameters: given a sample $\mathcal{X} = \{\mathbf{x}_n\}_{n=1}^N$: $\max_{\Theta} \mathcal{L}(\Theta; \mathcal{X}) = \sum_{n=1}^N \log p(\mathbf{x}_n; \Theta) = \sum_{n=1}^N \log \left(\sum_{k=1}^K p(\mathbf{x}|k)p(k)\right).$
- \mathcal{L} cannot be maximized in closed form over Θ ; it needs an iterative optimization algorithm. Many such algorithms exist (such as gradient descent), but there is a specially convenient one for mixture models (and more generally, for maximum likelihood with missing data).
- Expectation-Maximization (EM) algorithm: for Gaussian mixtures:
 - *E step*: given the current parameter values Θ , compute the posterior probability of component k given data point \mathbf{x}_n (for each k = 1, ..., K and n = 1, ..., N):

$$z_{nk} = p(k|\mathbf{x}_{n}; \boldsymbol{\Theta}) = \frac{p(\mathbf{x}_{n}|k)p(k)}{p(\mathbf{x}_{n}; \boldsymbol{\Theta})} = \frac{\pi_{k}|2\pi\boldsymbol{\Sigma}_{k}|^{-1/2}\exp\left(-\frac{1}{2}(\mathbf{x}_{n}-\boldsymbol{\mu}_{k})^{T}\boldsymbol{\Sigma}_{k}^{-1}(\mathbf{x}_{n}-\boldsymbol{\mu}_{k})\right)}{\sum_{k'=1}^{K}\pi_{k'}|2\pi\boldsymbol{\Sigma}_{k'}|^{-1/2}\exp\left(-\frac{1}{2}(\mathbf{x}_{n}-\boldsymbol{\mu}_{k'})^{T}\boldsymbol{\Sigma}_{k'}^{-1}(\mathbf{x}_{n}-\boldsymbol{\mu}_{k'})\right)} \in (0, 1).$$

- *M* step: given the posterior probabilities, estimate the parameters Θ : for $k = 1, \ldots, K$:

$$\pi_k = \frac{1}{N} \sum_{n=1}^N z_{nk} \qquad \mu_k = \frac{\sum_{n=1}^N z_{nk} \mathbf{x}_n}{\sum_{n=1}^N z_{nk}} \qquad \Sigma_k = \frac{\sum_{n=1}^N z_{nk} (\mathbf{x}_n - \boldsymbol{\mu}_k) (\mathbf{x}_n - \boldsymbol{\mu}_k)^T}{\sum_{n=1}^N z_{nk}}.$$

Similar to k-means, where the assignment and centroid steps correspond to the E and M steps. But in EM the assignments are soft $(z_{nk} \in [0, 1])$, while in k-means they are hard $(z_{nk} \in \{0, 1\})$.

- If we knew which component \mathbf{x}_n came from for each n = 1, ..., N, we'd not need the E step: a single M step that estimates each component's parameters on its set of points would suffice. This was the case in classification (where $\mathbf{x}|C_k$ is Gaussian): we were given (\mathbf{x}_n, y_n) .
- Each EM step increases \mathcal{L} or leaves it unchanged, but it takes an infinite number of iterations to converge. In practice, we stop when the parameters don't change much, or when the number of iterations reaches a limit.
- EM converges to a local optimum that depends on the initial value of Θ . Usually from *k*-means[?].
- User parameter: number of clusters K. Output: posterior probabilities $\{p(k|\mathbf{x}_n)\}$ and $\{\pi_k, \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k\}_{k=1}^K$.
- Parametric clustering: K clusters, assumed Gaussian.
- The fundamental advantage of Gaussian mixtures over k-means for clustering is that we can model the uncertainty in the assignments (particularly useful for points near cluster boundaries), and the clusters can be elliptical and have different proportions.

	k-means	EM for Gaussian mixtures
assignments z_{nk}	hard	soft, $p(k \mathbf{x}_n)$
probability model?	no	yes
number of iterations	finite	infinite
parameters	centroids $\{\boldsymbol{\mu}_k\}_{k=1}^K$	$\{\pi_k, oldsymbol{\mu}_k, oldsymbol{\Sigma}_k\}_{k=1}^K$